## **Spatial meshing** and **simplified manifold preconditioning** for scaling models of multivariate multi-type spatial data

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## Large scale multivariate spatial and spatiotemporal data



## **Species' occurrence & Communities**

- Misalignment?



• Spatial niches and shared resources • Highly multivariate, non-Gaussian data





### Multivariate spatial model with latent effects

Suppose we observe data at *n* locations (coordinates / pixels / voxels)

Probability model for outcome j  $y_{j}$ 

Linear predictor  $\eta_j$ 

example: Gaussian outcome



example: multivariate GLM with latent spatial factors

example: multi-species N-mixture models for large scale spatial abundance data

$$\ell_j(\boldsymbol{\ell}) \mid \eta_j(\boldsymbol{\ell}), \tau_j \sim P_j(\eta_j(\boldsymbol{\ell}), \tau_j)$$
$$\boldsymbol{\ell} \in \mathcal{D} \subset \Re^d \quad j = 1, \dots, q$$

$$(\boldsymbol{\ell}) = \boldsymbol{x}_j(\boldsymbol{\ell})^\top \boldsymbol{\beta}_j + w_j(\boldsymbol{\ell})$$

features or observed covariates latent/random/unobserved effects for dependent data

$$= \eta_j(\boldsymbol{\ell}) + \varepsilon_j(\boldsymbol{\ell}) \quad \varepsilon_j(\boldsymbol{\ell}) \sim N(0, \tau_j^2)$$



## Multivariate spatial model with latent effects

Suppose we observe data at *n* locations (coordinates / pixels / voxels)

Probability model for outcome *j*  $\mathcal{Y}$ 

> Linear predictor  $\eta_j$

example: Gaussian outcome

vector of random effects at location  $\ell$ 

**Gaussian Process (GP)** prior for  $w(\ell)$  at **any** set of locations  $L = \{\ell_1, \ldots, \ell_n\}$ 

$$\ell_j(\boldsymbol{\ell}) \mid \eta_j(\boldsymbol{\ell}), \tau_j \sim P_j(\eta_j(\boldsymbol{\ell}), \tau_j)$$
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$$(\boldsymbol{\ell}) = \boldsymbol{x}_j(\boldsymbol{\ell})^\top \boldsymbol{\beta}_j + w_j(\boldsymbol{\ell})$$

features or observed covariates latent/random/unobserved effects for dependent data

$$y_{j}(\boldsymbol{\ell}) = \eta_{j}(\boldsymbol{\ell}) + \varepsilon_{j}(\boldsymbol{\ell}) \quad \varepsilon_{j}(\boldsymbol{\ell}) \sim N(0, \tau_{j}^{2})$$
  
tion  $\boldsymbol{\ell} \quad \boldsymbol{w}(\boldsymbol{\ell}) = \begin{bmatrix} w_{1}(\boldsymbol{\ell}) \\ \vdots \\ w_{q}(\boldsymbol{\ell}) \end{bmatrix}$ 





## Scaling computations with Gaussian Processes

### **Gaussian Predictive Process & Inducing points**

Quiñonero-Candela and Rasmussen, 2005; Snelson and Ghahramani, 2007; Banerjee et al. 2008; Banerjee et al. 2010; Guhaniyogi et al. 2011; Finley, Banerjee, and Gelfand 2012; Low et al., 2015; Ambikasaran et al., 2016; Huang and Sun, 2018; Geoga et al., 2020 Exploit data structure

*Gilboa et al., 2015; Moran and Wheeler, 2020; Loper et al., 2020* **Fixed Rank Kriging** 

Cressie and Johannesson 2008

#### **Multi-resolution approximations**

Gramacy and Lee 2008; Fox and Dunson 2012; Katzfuss 2017

#### **Covariance Tapering**

Furrer, Genton, and Nychka 2006; Kaufman, Schervish, and Nychka 2008; Bevilacqua et al., 2019

#### Independent partitioning

Sang and Huang 2012; Stein 2014

#### **Composite likelihood**

Bai et al., 2012; Eidsvik et al., 2014; Bevilacqua and Gaetan, 2015

#### **Gaussian Random Markov Fields**

Cressie 1993; Rue 2001; Rue and Held 2005

#### **Vecchia's approximation & extensions**

Vecchia 1988; Stein et al. 2014; Gramacy and Apley 2015; Datta et al. 2016; Guinness 2018; Heaton et al. 2019; Katzfuss and Guinness 2019; Quiroz et al., 2019; Schafer et al. 2021



## $y_j(\boldsymbol{\ell}) \mid \eta_j(\boldsymbol{\ell}), \tau_j \sim P_j(\eta_j(\boldsymbol{\ell}), \tau_j)$ Probability model for outcome *j* $\ell \in \mathcal{D} \subset \Re^d \quad j = 1, \dots, q$ Linear predictor $\eta_j(\boldsymbol{\ell}) = \boldsymbol{x}_j(\boldsymbol{\ell})^\top \boldsymbol{\beta}_j + w_j(\boldsymbol{\ell})$ Prior for spatial random effects $m{w}(\cdot) \sim \mathrm{meshed} GP(m{0}, m{C}_{m{ heta}}, m{\mathcal{G}}, m{T})$







### **Data locations**

-	





### **Partition the domain**

1		
	 -	
	 -	
	 -	
1		





## Link partitions via DAG

-	-	
_		
	-	
	-	
	-	



 $y_j(\boldsymbol{\ell}) \mid \eta_j(\boldsymbol{\ell}), \tau_j \sim P_j(\eta_j(\boldsymbol{\ell}), \tau_j)$ Probability model for outcome *j*  $\ell \in \mathcal{D} \subset \Re^d \quad j = 1, \dots, q$ Linear predictor  $\eta_j(\boldsymbol{\ell}) = \boldsymbol{x}_j(\boldsymbol{\ell})^\top \boldsymbol{\beta}_j + w_j(\boldsymbol{\ell})$ Prior for spatial random effects  $m{w}(\cdot) \sim \mathrm{meshed}GP(m{0}, m{C}_{m{ heta}}, m{\mathcal{G}}, m{T})$ 

**Domain partitioning** 



Suppose we have *M* partitions: M $p(\boldsymbol{w}) = \prod p(\boldsymbol{w})$ j=1



M. Peruzzi, S. Banerjee & A.O. Finley (2022) Highly Scalable Bayesian Geostatistical Modeling via Meshed Gaussian Processes on Partitoned Domains. Journal of the American Statistical Association 117(538): 969-982. https://www.tandfonline.com/doi/full/10.1080/01621459.2020.1833889

#### **Sparse pattern DAG (mesh)**



$$(oldsymbol{w}_j \mid oldsymbol{w}_{[j]})$$



## Meshed Gaussian processes: Gibbs sampler

 $y_j(\boldsymbol{\ell}) \mid \eta_j(\boldsymbol{\ell}), \tau_j \sim P_j(\eta_j(\boldsymbol{\ell}), \tau_j)$ Probability model for outcome *j*  $\boldsymbol{\ell} \in \mathcal{D} \subset \Re^d \quad j = 1, \dots, q$ Linear predictor  $\eta_j(\boldsymbol{\ell}) = \boldsymbol{x}_j(\boldsymbol{\ell})^\top \boldsymbol{\beta}_j + w_j(\boldsymbol{\ell})$ Prior for spatial random effects  $m{w}(\cdot) \sim \mathrm{meshed} GP(m{0}, m{C}_{m{ heta}}, m{\mathcal{G}}, m{T})$ 





M. Peruzzi, S. Banerjee & A.O. Finley (2022) Highly Scalable Bayesian Geostatistical Modeling via Meshed Gaussian Processes on Partitoned Domains. Journal of the American Statistical Association 117(538): 969-982. https://www.tandfonline.com/doi/full/10.1080/01621459.2020.1833889

#### Meshed GP:

computationally **cheap** because

$$\begin{array}{l} \left[ \boldsymbol{\theta} \mid \boldsymbol{w} \right] \propto \prod_{j=1}^{M} p(\boldsymbol{w}_{j} \mid \boldsymbol{w}_{[j]}, \boldsymbol{\theta}) p(\boldsymbol{\theta}) \\ & \text{id} \\ \left[ \boldsymbol{w}_{j} \mid \boldsymbol{w}_{[j]}, \boldsymbol{\theta} \right] = N(\boldsymbol{w}_{j}; \boldsymbol{H}_{j} \boldsymbol{w}_{[j]}, \boldsymbol{R}_{j}) \\ & \text{here} \\ & \boldsymbol{f}_{j} = \boldsymbol{C}_{j,[j]} \boldsymbol{C}_{[j]}^{-1} \\ & \boldsymbol{R}_{j} = \boldsymbol{C}_{j,j} - \boldsymbol{H}_{j} \boldsymbol{C}_{[j],j} \\ & \text{id} \ \boldsymbol{C}_{[j]}^{-1} \text{ is } \boldsymbol{small for all } \boldsymbol{j}. \end{array}$$



## Serengeti NDVI data: gap filling via Meshed GP

- Normalized Difference Vegetation Index (NDVI) measured by LANDSAT over Serengeti park area
- 16M locations in space & time
- Cloud cover obfuscates images
- Posterior sampling for a **bivariate latent MGP**
- Results in less than 2 days

**Observed NDVI** 



Predicted NDVI





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### • Model univariate NDVI outcome using spatiotemporally varying coefficient regression on 2 inputs (intercept, elevation)

#### Residual spatial effect



Covariance with estim



#### Elevation: effect on NDV



Elevation: nonzero effect locations (95% cred. int.)





## **MODIS / GHCN multi-modality data with SPAMTREES**

- 5 outcomes from 2 different sources
- MODIS **satellite** data
- GHCN land-based station data
- PRCP more sparsely observed
- Misalignment
- Data size ~ 1M
- Compute time 16 hours or less







M. Peruzzi & D.B. Dunson (2022) Spatial Multivariate Trees for Big Data Bayesian Regression. Journal of Machine Learning Research 23(17):1–40. https://www.jmlr.org/papers/v23/20-1361.html





## NASA LiDAR data over the Tanana forest in Alaska



#### **GriPS**: a method that involves knot grid & parameter expansion for improved efficiency in estimating Matérn models

#### Application

- Tanana forest, Alaska: 2 outcomes at **2.5M locations**
- Forest cover and canopy height (measured via LiDAR)
- Awkward measurement pattern
- Highly customized setup for MGPs for under 48h compute time


M. Peruzzi, S. Banerjee, D.B. Dunson & A.O. Finley (2021) Grid-Parametrize-Split (GriPS) for Improved Scalable Inference in Spatial Big Data Analysis. https://arxiv.org/abs/2101.03579





Width of 95% d 2 5 10 3 Width of 95% c. 0.25 0.50 0.75





## **Simplified manifold preconditioning adaptation** for sampling with multivariate non-Gaussian outcomes



**M. Peruzzi** & D.B. Dunson (2022) Spatial Meshing for General Bayesian Multivariate Models. *https://arxiv.org/abs/2201.10080* 



### Meshed Gaussian processes: MCMC with non-Gaussian outcomes

 $y_j(\boldsymbol{\ell}) \mid \eta_j(\boldsymbol{\ell}), \tau_j \sim P_j(\eta_j(\boldsymbol{\ell}), \tau_j)$ Probability model for outcome *j*  $\ell \in \mathcal{D} \subset \Re^d \quad j = 1, \dots, q$ Linear predictor  $\eta_j(\boldsymbol{\ell}) = \boldsymbol{x}_j(\boldsymbol{\ell})^\top \boldsymbol{\beta}_j + w_j(\boldsymbol{\ell})$ Prior for spatial random effects  $m{w}(\cdot) \sim \mathrm{meshed} GP(m{0}, m{C}_{m{ heta}}, m{\mathcal{G}}, m{T})$ 













• Propose move from  $w_j$  to  $w_j^*$  via the function  $q(\cdot \mid w_j)$ • Target density is the full conditional distribution of  $w_j$  $p(w_j \mid {}_{everything}) \propto p(y_j \mid w_j,.$ 

$$\dots)\pi(w_j \mid w_{[j]}) \prod_{j \to i} \pi(w_i \mid w_j, w_{[i]\setminus j})$$

- Propose move from  $w_j$  to  $w_j^*$  via the function  $q(\cdot \mid w_j)$
- Target density is the full conditional distribution of  $w_j$

$$p(w_j \mid {}_{everything}) \propto p(y_j \mid w_j, \dots) \pi(w_j \mid w_{[j]}) \prod_{j \to i} \pi(w_i \mid w_j, w_{[i]\setminus j})$$

• Accept based on the ratio

$$\alpha = \frac{p(w_j^* \mid everything)q(w_j \mid w_j^*)}{p(w_j \mid everything)q(w_j^* \mid w_j)}$$

• If we can sample from full conditional itself, then set

$$q(w_j \mid w_j^*) = p(w_j \mid \text{everything})$$

to get  $\alpha = 1$  which leads to a Gibbs sampler

• If we cannot sample from full conditional itself, then what proposal do we use?



### **Preconditioned MALA**

 $w_j^* \sim N(w_j + \frac{\varepsilon^2}{2}Gg, \varepsilon^2 G)$ 



and G is a preconditioner matrix which we need to specify

• Use accepted values to build G based on the Markov chain's past? **Slow to adapt** • Use a position-dependent  $G_{w_j}$  à la Riemann manifold MALA? Expensive update

 $g = \nabla p(w_j \mid everything)$ 

gradient of the target density



## Simplified manifold preconditioner adaptation (SiMPA)

at iteration m of our Metropolis algorithm:

 $w_j^* \sim N(w_j + \frac{\varepsilon^2}{2}G_m g, \varepsilon^2 G_m)$ 

• accept or reject as usual based on this proposal • afterwards, with probability  $\gamma_m \downarrow 0$  set

$$G_m^{-1} = G_{m-1}^{-1} + \kappa (\nabla \nabla$$

• Asymptotically constant preconditioner: cheap update • Uses target density's second order information, similar to Riemannian manifold MALA: efficient update

 $7p(w_j \mid everything)) - G_{m-1}^{-1})$ 



## **SiMPA**: Simplified manifold preconditioner adaptation – simulated data

### Poisson outcome 1

Outcome 1: Outcome 1: 0 25 50 75 -5.0 -2.5 0.0 2.5 Log-intensity full data 1.0 0.5 -Northing 0.0 -0.5

0.0 Easting 0.5 0.0 Easting

### Poisson outcome 2



#### QMGP-SiMPA



- Two related count outcomes
  - *n=3600*
  - leave out 20% of data
- Goals
  - prediction
  - recovery of latent log-intensity
  - uncertainty quant. about log-intensity
- meshed GPs with SiMPA
  - results in <10s
  - orders of magnitude faster than full GP
- using R package **meshed** 0.2.1 (it's on CRAN)



## **SiMPA**: Simplified manifold preconditioner adaptation. Simulated data





R package meshed: demonstrating MGPs and SiMPA on the **North American Breeding Bird survey** 



## North American Breeding Bird Survey data

- Somewhat small number of spatial locations (~ 4100)
- Large number of binary outcomes, q=27
- Effective data size ~100k
- Presence/Absence data about birds of the *Passeriforme* order
- Create a **test set** using 20% of data
- Target: classification performance for all bird species measured via area under ROC curve (AUC)
- **Compare** with popular software **Hmsc** (non-spatial, spatial w/ NNGP) commonly used by biologists







## **Relevant articles**

- Peruzzi M, Dunson DB (2023) Spatial meshing for general multivariate Bayesian models (arXiv: 2201.10080)
- Peruzzi M, Banerjee S, Finley AO (2022) Meshed GPs. JASA 117(538): 969–982. (doi: 10.1080/01621459.2020.1833889)
- Peruzzi M, Dunson DB (2022) Spatial Multivariate Trees. JMLR (https://www.jmlr.org/papers/v23/20-1361.html)
- Peruzzi M, Banerjee S, Dunson DB, Finley AO (2021) Grid-Parametrize-Split (GriPS) (submitted, arXiv: 2101.03579)
- Jin B, Peruzzi M, Dunson DB (2021) Bags of DAGs (submitted, arXiv: 2112.11870)

#### **R** PACKAGES

- **meshed** R package: *https://CRAN.R-project.org/package=meshed*
- **spamtree** R package: *https://CRAN.R-project.org/package=spamtree*
- **spammix** *R* package: https://github.com/mkln/spammix (multi-species N-mixture modeling with latent spatial factors)

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#### Thank you!

